

# Structural VARs

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## (Lack of) Identification

- So far, we considered reduced form VARs, say,

$$y_t = \Phi_1 y_{t-1} + u_t, \quad \mathbb{E}[u_t u_t'] = \Sigma \quad (1)$$

- Error terms  $u_t$  have the interpretation of one-step ahead forecast errors.
- If the eigenvalues of  $\Phi_1$  are inside the unit-circle then  $y_t$  has the following moving-average (MA) representation in terms of  $u_t$ :

$$y_t = (I - \Phi_1 L)^{-1} u_t = \sum_{j=0}^{\infty} \Phi_1^j u_{t-j} = \sum_{j=0}^{\infty} C_j u_{t-j} \quad (2)$$

- DSGE models suggest that the one-step ahead forecast errors are functions of some fundamental shocks, such as technology shocks, preference shocks, or monetary policy shocks.

## (Lack of) Identification

- Let  $\epsilon_t$  a vector of such fundamental shocks and assume that  $\mathbf{E}[\epsilon_t \epsilon_t'] = \mathcal{I}$ . Moreover, assume that

$$u_t = \Phi_\epsilon \epsilon_t. \quad (3)$$

- Then we can express the VAR in structural form as follows

$$y_t = \Phi_1 y_{t-1} + \Phi_\epsilon \epsilon_t \quad (4)$$

$$\Phi_\epsilon^{-1} y_t = \Phi_\epsilon^{-1} \Phi_1 y_{t-1} + \epsilon_t$$

- The moving-average representation of  $y_t$  in terms of the structural shocks is given by

$$y_t = \sum_{j=0}^{\infty} \Phi_1^j \Phi_\epsilon \epsilon_{t-j} = \sum_{j=0}^{\infty} C_j \Phi_\epsilon \epsilon_{t-j}. \quad (5)$$

## (Lack of) Identification

- For (1) and (4) the matrix  $\Phi_\epsilon$  has to satisfy the restriction

$$\Phi_\epsilon \Phi_\epsilon' = \Sigma \tag{6}$$

Notice that the matrix  $\Phi_\epsilon$  has  $n^2$  elements.

- The covariance relationship, unfortunately, generates only  $n(n+1)/2$  restrictions and does not uniquely determine  $\Phi_\epsilon$ .
- This creates an identification problem since all we can estimate from the data is  $\Phi_1$  and  $\Sigma$ .

## (Lack of) Identification

- In order to make statements about the propagation of structural shocks  $\epsilon_t$  we have to make further assumptions. The papers (see course outline) by Cochrane (1994), Christiano and Eichenbaum (1999), and Stock and Watson (2001) survey such identifying assumptions. A cynical view of this literature is the following:
  1. Propose an identification scheme, that determines all elements of  $\Phi_\epsilon$ .
  2. Compute impulse response functions.
  3. If impulse response functions are plausible, then stop; else, declare a “puzzle” and return to 1.

## (Lack of) Identification

- Here are some famous “puzzles:”
  1. “Liquidity Puzzle:” When identifying monetary policy shocks as surprise changes in the stock of money one often finds that interest rates fall when the money stock is lowered.
  2. “Price Puzzle:” When identifying monetary policy shocks as surprise changes in the Federal Funds Rate, one often finds that prices fall after a drop in interest rates.
- These “puzzles” are typically resolved by considering more elaborate identification schemes.

## Identification Schemes

- We begin by decomposing the covariance matrix into the product of lower triangular matrices (Cholesky Decomposition):

$$\Sigma = AA', \quad (7)$$

where  $A$  is lower triangular. If  $\Sigma$  is non-singular the decomposition is unique.

- Let  $\Omega$  be an orthonormal matrix, meaning that  $\Omega\Omega' = \Omega'\Omega = \mathcal{I}$ .
- We can characterize the relationship between the reduced form and the structural shocks as follows

$$u_t = A\Omega\epsilon_t \quad (8)$$

- Notice that

$$\mathbf{IE}[u_t u_t'] = \mathbf{IE}[A\Omega\epsilon_t \epsilon_t' \Omega' A'] = A\Omega \mathbf{IE}[\epsilon_t \epsilon_t'] \Omega' A' = A\Omega\Omega' A' = AA' = \Sigma. \quad (9)$$

## Identification Schemes

- In general, it is quite tedious to characterize the space of orthonormal matrices. Let's try for  $n = 2$ :

$$\Omega(\varphi) = \begin{bmatrix} \cos \varphi & -\sin \varphi \\ \sin \varphi & \cos \varphi \end{bmatrix} \quad (10)$$

where  $\varphi \in (-\pi, \pi]$ .

- Notice that, for instance,

$$\Omega(\pi/2) = -\Omega(-\pi/2) \quad (11)$$

which means that only the signs of the impulse responses change but not the shape.

- Identification schemes impose restrictions on  $\varphi$ .

## Short-run Restrictions

- Suppose that

$$y_t = \begin{bmatrix} \text{Fed Funds Rate} \\ \text{Output Growth} \end{bmatrix}, \quad \epsilon_t = \begin{bmatrix} \epsilon_{R,t} \\ \epsilon_{z,t} \end{bmatrix} = \begin{bmatrix} \text{Monetary Policy Shock} \\ \text{Technology Shock} \end{bmatrix}.$$

Moreover, we assume that the central bank does not react contemporaneously to technology shocks because data on aggregate output only become available with a one-quarter lag.

This assumption can be formalized through  $\varphi = 0$ . Then

$$u_t = \begin{bmatrix} a_{11} & 0 \\ a_{21} & a_{22} \end{bmatrix} \begin{bmatrix} \epsilon_{R,t} \\ \epsilon_{z,t} \end{bmatrix}. \quad (12)$$

Reference: Sims (1980).

## Long-run Restrictions

- Now suppose that

$$y_t = \begin{bmatrix} \text{Inflation} \\ \text{Output Growth} \end{bmatrix}, \quad \epsilon_t = \begin{bmatrix} \epsilon_{R,t} \\ \epsilon_{z,t} \end{bmatrix} = \begin{bmatrix} \text{Monetary Policy Shock} \\ \text{Technology Shock} \end{bmatrix}$$

Moreover,

$$y_t = \left( \sum_{j=0}^{\infty} C_j L^j \right) u_t = C(L) u_t. \quad (13)$$

Consider the following assumption: monetary policy shocks do not raise output in the long-run.

## Long-run Restrictions

- Let's examine the moving average representation of  $y_t$  in terms of the structural shocks

$$\begin{aligned}
 y_t &= \begin{bmatrix} c_{11}(L) & c_{12}(L) \\ c_{21}(L) & c_{22}(L) \end{bmatrix} \begin{bmatrix} a_{11} & 0 \\ a_{21} & a_{22} \end{bmatrix} \begin{bmatrix} \cos \varphi & -\sin \varphi \\ \sin \varphi & \cos \varphi \end{bmatrix} \begin{bmatrix} \epsilon_{R,t} \\ \epsilon_{z,t} \end{bmatrix} \\
 &= \begin{bmatrix} \cdot & \cdot \\ a_{11} \cos \varphi c_{21}(L) + (a_{21} \cos \varphi + a_{22} \sin \varphi) c_{22}(L) & \cdot \end{bmatrix} \begin{bmatrix} \epsilon_{R,t} \\ \epsilon_{z,t} \end{bmatrix} \\
 &= \begin{bmatrix} d_{11}(L) & d_{12}(L) \\ d_{21}(L) & d_{22}(L) \end{bmatrix} \begin{bmatrix} \epsilon_{R,t} \\ \epsilon_{z,t} \end{bmatrix}
 \end{aligned}$$

- Suppose that in period  $t = 0$  log output and log prices are equal to zero. Then the log-level of output and prices in period  $t = T > 0$  is given by

$$y_T^c = \sum_{t=1}^T y_t = \sum_{t=1}^T \sum_{j=0}^{\infty} D_j \epsilon_{t-j} \quad (14)$$

## Long-run Restrictions

- Now consider the derivative

$$\frac{\partial y_T^c}{\partial \epsilon_1'} = \sum_{j=0}^{T-1} D_j \quad (15)$$

- Letting  $T \rightarrow \infty$  gives us the long-run response of the level of prices and output to the shock  $\epsilon_1$ :

$$\frac{\partial y_\infty^c}{\partial \epsilon_1'} = \sum_{j=0}^{\infty} D_j = D(1) \quad (16)$$

- Here, we want to restrict the long-run effect of monetary policy shocks on output:

$$d_{21}(1) = 0 \quad (17)$$

## Long-run Restrictions

- This leads us to the equation

$$[a_{11}c_{21}(1) + a_{21}c_{22}(1)] \cos \varphi + a_{22}c_{22}(1) \sin \varphi = 0. \quad (18)$$

- Notice that the equation has two solutions for  $\varphi \in (-\pi, \pi]$ . Under one solution a positive monetary policy shock is contractionary, under the other solution it is expansionary. The shape of the responses is, of course, the same.
- Reference: Blanchard and Quah (1989).

## Sign Restrictions

- Again consider

$$y_t = \begin{bmatrix} \text{Inflation} \\ \text{Output Growth} \end{bmatrix}, \quad \epsilon_t = \begin{bmatrix} \epsilon_{R,t} \\ \epsilon_{z,t} \end{bmatrix} = \begin{bmatrix} \text{Monetary Policy Shock} \\ \text{Technology Shock} \end{bmatrix}$$

- Our identification assumption is: upon impact, a monetary policy shock raises both prices and output. It can be verified that

$$\frac{\partial y_t}{\partial \epsilon_{R,t}} = \begin{bmatrix} a_{11} \cos \varphi + (a_{21} \cos \varphi + a_{22} \sin \varphi) \\ a_{11} \cos \varphi + (a_{21} \cos \varphi + a_{22} \sin \varphi) \end{bmatrix}. \quad (19)$$

## Sign Restrictions

- Thus, we obtain the sign restrictions

$$0 < a_{11} \cos \varphi + (a_{21} \cos \varphi + a_{22} \sin \varphi)$$

$$0 < a_{11} \cos \varphi + (a_{21} \cos \varphi + a_{22} \sin \varphi)$$

which restrict  $\varphi$  to be in a certain subset of  $(-\pi, \pi]$  and will generate a range of responses.

- References: Canova and De Nicolò (2002), Faust (1998), Uhlig (2005).

# Impulse Responses and Variance Decompositions

- Impulse responses are defined as

$$\frac{\partial y_{t+h}}{\partial \epsilon'_t} = C_h \Phi_\epsilon \quad (20)$$

and correspond to the MA coefficient matrices in the moving average representation of  $y_t$  in terms of structural shocks.

# Impulse Responses and Variance Decompositions

- The covariance matrix of  $y_t$  is given by

$$\Gamma_{yy,0} = \sum_{j=0}^{\infty} C_j \Phi_{\epsilon} \mathcal{I} \Phi_{\epsilon}' C_j' \quad (21)$$

Let  $\mathcal{I}^i$  be matrix for which element  $i, i$  is equal to one and all other elements are equal to zero. Then we can define the contribution of the  $i$ 'th structural shock to the variance of  $y_t$  as

$$\Gamma_{yy,0}^{(i)} = \sum_{j=0}^{\infty} C_j \Phi_{\epsilon} \mathcal{I}^{(i)} \Phi_{\epsilon}' C_j' \quad (22)$$

Thus the fraction of the variance of  $y_{l,t}$  explained by shock  $i$  is

$$\frac{[\Gamma_{yy,0}^{(i)}]_{ll}}{[\Gamma_{yy,0}]_{ll}}. \quad (23)$$

## Alternative Setups

- Sims and Zha (1998) and subsequent work start out from the specification

$$A_0 y_t = A_1 y_{t-1} + \dots + A_p y_{t-p} + \epsilon_t$$

where  $\epsilon_t$ 's are structural shocks.

- Analyze the model directly in terms of  $A$  matrices.
- Impose identification restrictions on  $A_0$ .

## References

- Blanchard, Oliver Jean, and Danny Quah (1989): “The Dynamic Effects of Aggregate Demand and Supply Disturbances,” *American Economic Review*, **79**, 655-673.
- Canova, F. and G. De Nicoló, “Monetary Disturbances Matter for Business Cycle Fluctuations in the G-7.” *Journal of Monetary Economics* **49**(6) (2002), 1131-1159.
- Cochrane, John H. (1994): “Shocks,” *Carnegie-Rochester Conference Series on Public Policy*, **41**, 295-364.
- Doan, Thomas, Robert Litterman, and Christopher Sims (1984): “Forecasting and Conditional Projections Using Realistic Prior Distributions,” *Econometric Reviews*, **3**, 1-100.
- Faust, Jon (1998): “The Robustness of Identified VAR Conclusions about Money,”

*Carnegie Rochester Conference Series*, **49**, 207-244.

- Sims, Christopher, and Tao Zha (1998): “Bayesian Methods for Dynamic Multivariate Models,” *International Economic Review*, **39**, 949-968.
- Stock, James J. and Mark W. Watson (2001): “Vector Autoregressions,” *Journal of Economic Perspectives*, **15**, 101-115.
- Uhlig, Harald (2005) “What are the Effects of Monetary Policy? Evidence from an Agnostic Identification Procedure,” *Journal of Monetary Economics*.